Solutions 6

Exercise 11.7

For Bayesian linear model, MMSE estimation is identical to MAP estimation since $p(\boldsymbol{\theta}|\mathbf{x})$ is Gaussian. But MAP estimation maximizes $p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})$ with no prior information, equivalent to maximizing $p(\mathbf{x}|\boldsymbol{\theta})$. In the Bayesian model, $p(\mathbf{x}|\boldsymbol{\theta}) = p(\mathbf{x};\boldsymbol{\theta})$. Thus, maximizing $p(\mathbf{x};\boldsymbol{\theta})$, which yields the MLE or MVUE, also yields the MMSE. Exercise 11.11

$$R = \mathbb{E}[C(\boldsymbol{\epsilon})]$$

$$= \int \int C(\boldsymbol{\epsilon}) p(\mathbf{x}, \boldsymbol{\theta}) d\mathbf{x} d\boldsymbol{\theta}$$

$$= \int \left(\int C(\boldsymbol{\epsilon}) p(\boldsymbol{\theta} | \mathbf{x}) d\boldsymbol{\theta} \right) p(\mathbf{x}) d\mathbf{x}$$

$$\int C(\boldsymbol{\epsilon}) p(\boldsymbol{\theta} | \mathbf{x}) d\boldsymbol{\theta} = \int_{||\boldsymbol{\epsilon}|| > \delta} p(\boldsymbol{\theta} | \mathbf{x}) d\boldsymbol{\theta} = 1 - \int_{||\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}|| < \delta} p(\boldsymbol{\theta} | \mathbf{x}) d\boldsymbol{\theta}$$

as $\delta \to 0$, we minimize the above by choosing $\hat{\theta} = \arg \max_{\theta} p(\theta | \mathbf{x})$. Exercise 11.12

If $\boldsymbol{\alpha} = \mathbf{A}\boldsymbol{\theta}$, then $\partial \boldsymbol{\alpha} / \partial \boldsymbol{\theta} = \mathbf{A}$

$$p(\mathbf{x}, \boldsymbol{\alpha}) = \frac{p(\mathbf{x}, \boldsymbol{\theta})}{|\det \frac{\partial \boldsymbol{\alpha}/}{\partial \boldsymbol{\theta}}|} = \frac{p(\mathbf{x}, \boldsymbol{\theta})}{|\det \mathbf{A}|}$$

However, **A** does not depend on $\boldsymbol{\alpha}$ and $\boldsymbol{\theta} = \mathbf{A}^{-1}\boldsymbol{\alpha}$, so that

$$p(\mathbf{x}, \boldsymbol{\alpha}) = \frac{p_{x,\theta}(\mathbf{x}, \mathbf{A}^{-1}\boldsymbol{\alpha})}{|\det \mathbf{A}|}$$

The MAP estimator of $\boldsymbol{\alpha}$ maximizes $p_{x,\theta}(\mathbf{x}, \mathbf{A}^{-1}\boldsymbol{\alpha})$, equivalent to maximizing $p(\mathbf{x}, \boldsymbol{\theta})$ because $\boldsymbol{\theta} = \mathbf{A}^{-1}\boldsymbol{\alpha}$ is invertible. Thus, $\hat{\boldsymbol{\alpha}} = \mathbf{A}\hat{\boldsymbol{\theta}}$

Exercise 12.2

From (12.27) in page 391, we can get

$$\hat{A} = \mu_A + \left(\frac{1}{\sigma_A^2} + \frac{\mathbf{h}^T \mathbf{h}}{\sigma^2}\right)^{-1} \frac{\mathbf{h}^T}{\sigma^2}$$

where $\mathbf{h} = [1, r, ..., r^{N-1}]^T$. Thus,

$$\hat{A} = \mu_A + \frac{\sum_{n=0}^{N-1} r^n (x[n] - r^n \mu_A)}{\frac{\sigma^2}{\sigma_A^2} + \sum_{n=0}^{N-1} r^{2n}}$$

From (12.29) and (12.30), we get

Bmse
$$(\hat{A}) = \frac{1}{\frac{1}{\sigma_A^2} + \frac{1}{\sigma^2} \sum_{n=0}^{N-1} r^{2n}}$$

Exercise 12.14

To minimize $\mathbb{E}[(x[n] - \hat{x}[n])^2]$, we use the orthogonality principle, i.e.

$$\mathbb{E}[(x[n] - \hat{x}[n])x[n-l]] = 0, \qquad l = -M, ..., M(l \neq 0)$$
$$r_{xx}(l) = \mathbb{E}[\sum_{k} a_k x[n-k]x[n-l]] = \sum_{k} a_k r_{xx}(l-k)$$

To show that $a_{-k} = a_k$, we let k' = -k

$$r_{xx}(l) = \sum_{k'=-M, k'\neq 0}^{M} a_{-k'} r_{xx}(l+k')$$

Let l' = -l

$$r_{xx}(-l') = \sum_{k'=-M, k'\neq 0}^{M} a_{-k'} r_{xx}(-l'+k')$$
$$r_{xx}(l') = \sum_{k'=-M, k'\neq 0}^{M} a_{-k'} r_{xx}(l'-k')$$

Hence $r_{xx}(-k) = r_{xx}(k)$. But these are the same set of equation for which there is a unique solution. Hence $a_{-k} = a_k$. This must be true since the correlation of x[n] with x[n+k] is the same as that with x[n-k], due to the even symmetry. Exercise 12.19

$$\hat{x}[n] = \sum_{k=1}^{N} h(k)x[n-k]$$
$$\mathbb{E}[(x[n] - \hat{x}[n])x[n-l]] = 0$$
$$r_{xx}(l) = \sum_{k=1}^{N} h(k)\mathbb{E}(x[n-k]x[n-l]) = \sum_{k=1}^{N} h(k)r_{xx}(l-k)$$

The equations are independent of n since in deriving (12.65) we assumed n = N was the

index of the sample to be predicted. Hence the ACF does not depend on n

$$M_{\hat{x}} = \mathbb{E}[(x[n] - \hat{x}[n])x[n]] - \mathbb{E}[(x[n] - \hat{x}[n])\hat{x}[n]]$$

= $\mathbb{E}[x^{2}[n]] - \sum_{k=1}^{N} h(k)\mathbb{E}[x[n-k]x[n]]$
= $r_{xx}(0) - \sum_{k=1}^{N} h(k)r_{xx}(k)$

Exercise 12.20

From the previous problem, we have

$$r_{xx}(l) = \sum_{k=1}^{N} h(k) r_{xx}(l-k)$$

must be solved for the optimal one-step prediction. But for an AR(N) process, we know that

$$r_{xx}(l) = -\sum_{k=1}^{N} a[k]r_{xx}(l-k)$$

which are the Yale-Walker equations. Hence the solution for the h(k) is unique,

$$h(k) = -a[k]$$

so that $\hat{x}[n] = -\sum_{k=1}^{N} a[k]x[n-k]$ and the MMSE is

$$M_{\hat{x}} = r_{xx}(0) - \sum_{k=1}^{N} h(k) r_{xx}(k)$$

= $r_{xx}(0) + \sum_{k=1}^{N} a[k] r_{xx}(k)$
= σ_u^2